Genetic Susceptibility to Dental Caries on Pit and Fissure and Smooth Surfaces

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Heritability of caries scores for both PFS and SMS in the primary dentition was greater than in the permanent dentition and total dentition. With one exception, the genetic correlation between PFS and SMS caries scores was not significantly different from 100%, indicating that (mostly) common genes are involved in the risk of caries for both surface types. Genetic correlation for the primary dentition dfs (decay + filled surfaces) was significantly less than 100% (p < 0.001), indicating that genetic factors may exert differential effects on caries risk in PFS versus SMS in the primary dentition.

Dental caries is a multifactorial disease widely acknowledged to be caused by a combination of environmental and behavioral factors and genetic predispositions. The environmental risk factors for dental caries have been studied for decades, and include dietary behaviors, bacterial flora, transmission of bacteria among hosts, hygiene, salivary composition and flow rate, tooth positional and morphological features, and fluoride ex-
Dental Caries Genetics: Pit and Fissure and Smooth Surfaces

**Subjects and Methods**

**Recruitment and Data Collection**

The principal goal of COHRA was to identify the community-, family- and individual-level factors related to oral diseases in the Appalachian population [Polk et al., 2008], a group with notably poor oral health in comparison to the greater US population [Grembowski et al., 1987; Purnell and Counts, 1998; From the Centers for Disease Control and Prevention, 1999; Janes et al., 1999; Public Health and Aging, 2003]. Household-based recruitment of COHRA participants was carried out in Allegheny, Washington and McKean counties of Pennsylvania, and Webster and Nicholas counties of West Virginia. To meet eligibility criteria, households were required to contain one biological parent-offspring pair with the child being 1–18 years of age. All recruitment was conducted without regard to participants’ oral health status, and all members of eligible households were invited into the study without regard to biological and/or legal relationships. Participants were excluded according to the following criteria: neurological impairment, severe physical or mental disability, and psychosis. Households were excluded if one member of the biological parent-offspring pair suffered from diminished capacity to resist infection or from blood clots. Written informed consent was obtained from all adult participants. Assent with parent-guardian-written consent was obtained for all child participants. All forms and procedures were approved by the COHRA research committee and the Institutional Review Boards of the University of Pittsburgh and West Virginia University.

Our population-based household recruitment methods yielded a sample that accurately reflects the Appalachian population, including households from a range of socioeconomic statuses. However, as with the general Appalachian population, the majority of households recruited in this study are low-income (i.e. the median annual household income was less than $25,000), indicating that some oral health risk factors may be more prevalent in this population compared to the greater US, although we believe the effects of such risk factors may be largely generalizable.

732 households comprising a total of 2,663 individuals from 740 biological kinships were enrolled. Families consisting of confirmed biological relatives ranged from 1 to 20 individuals (mean = 4.72 members). In total, 3,232 relative pairs were available for analysis, which included 1,817 parent-offspring pairs, 756 sibling pairs, 347 half-sibling pairs, 120 avuncular (i.e. uncle/aunt-nephew/niece) pairs, 104 first-cousin pairs, and 88 other relative pairs. The accuracy of reported familial relationships was validated using panels of ancestry-informative [Marosy et al., 2007] and whole-genome [Cornelis et al., 2010] genetic marker data genotyped by the Center for Inherited Disease Research at Johns Hopkins University and quality checked jointly with the Coordinating Center for the NIH Genes and Environment Initiative (GENEVA) [Laurie et al., 2010]. Standard relationship-testing methods [O’Connell and Weeks, 1998] were used to identify and correct relationship errors prior to conducting the analysis reported herein.
Dental caries were assessed during dental exams conducted by dentists or research dental hygienists (the latter being calibrated to the reference dentist at least once per year). Inter- and intraexaminer concordance of caries assessments was high [Polk et al., 2008; Wendell et al., 2010]. Full details of the dental exams have been previously described [Polk et al., 2008; Wang et al., 2010]. In brief, teeth were scored according to NIH/NIDCR-approved methods for assessing dental caries for research purposes via visual inspection, with each surface being classified as sound, precavitated, decayed (according to a four-level classification system), filled, missing due to decay, hypoplastic, or missing due to reasons other than decay, in accordance with the World Health Organization DMFS/dfs scale. Two classes of tooth surfaces were defined: PFS included occlusal surfaces of molars and premolars, buccal surfaces of lower molars, and lingual surfaces of upper molars; SMS included all other tooth surfaces. Surface-level caries scores were used to generate the following four composite caries scores: (1) PFS DMFS score, which was the summation of decayed (D), missing due to decay (M), and filled (F) permanent dentition for PFS; (2) proportion PFS DMFS score, which was calculated by dividing the PFS DMFS score by the total number of permanent dentition PFS (i.e., surfaces present plus surfaces missing due to decay); (3) SMS DMFS score, which was the summation of D, M, and F permanent dentition SMS; (4) proportion SMS DMFS, which was the SMS DMFS score divided by the number of permanent dentition SMS for which we have data. Four analogous caries scores were similarly calculated for primary tooth surfaces (i.e., PFS dfs, proportion PFS dfs, SMS dfs, and proportion SMS dfs), with the exception that the missing tooth component was not included in primary dfs scores due to measurement issues introduced by unerupted primary teeth. Likewise, composite variables were calculated for ‘total’ dentition (DMFS + dfs, etc.) by considering all teeth present without distinguishing between permanent and primary dentitions. Proportion caries scores were analyzed in order to account for the differences in number of teeth present among participants. Based on findings from previous genetic analyses in this sample [Wang et al., 2010], we have included pre-cavitated lesions (i.e., white spots) in the composite caries scores. Permanent and primary caries scores were analyzed for individuals based solely on the presence of teeth of each dentition type (i.e., not limited to specific age ranges). Third molars were excluded when calculating caries scores in this study. Endodontulous family/household members, including infants, were excluded from the study (n = 60).

Statistical Analysis

The modeling framework used in this study estimates the heritability of a single phenotype and/or the genetic correlation between two phenotypes by comparing phenotype measurements across all pairs of relatives and nonrelatives while conditioning on the expected genetic sharing between pairs (e.g., that parents share 50% of their genetic material with their offspring, siblings share 50%, half-siblings share 25%, unrelated individuals share 0%, etc.). Heritability estimates are interpreted as the portion of phenotype variation attributable to the cumulative effect of all genes affecting the phenotype. Genetic correlation estimates are interpreted as the degree to which two phenotypes are influenced by the same genetic effects.

Analyses were performed using variance decomposition methods, which partition the trait variance into environmental, heritable, and residual error components, while conditioning on the familial relationships among the study participants. This model takes the general form:

$$y_i = \mu + \sum_j \beta_j x_{ij} + g_i + e_i,$$

where $y_i$ is the dental phenotype for the i-th individual, $\mu$ is the sample mean, $x_{ij}$ are the values of the predictors, $\beta_j$ are the regression coefficients for the j-th predictor, $g_i$ is the individual additive polygenic effect (based on expected genetic sharing due to the familial relationship between relative pairs), and $e_i$ is the individual residual error effect [Almasy and Blangero, 1998]. Pedigree-based likelihood methods were used to estimate model parameters, from which we calculated residual heritability (h^2, i.e. the proportion of phenotype variance attributable to genetics after adjustment for covariates, age and sex). The statistical significance of heritability was assessed by the likelihood ratio test, comparing the model where the polygenic effect was estimated to the model where the polygenic effect was constrained to be zero. This test statistic asymptotically follows 50:50 mixture of the $\chi^2$ distribution with 1 degree of freedom and a point mass at zero.

A bivariate extension of the above univariate model was used to estimate the genetic correlation between caries scores, which describes the extent to which family structure explains the covariance of the two traits (i.e., whether the traits vary together through the family, based on expected genetic sharing among relatives). This model can loosely be interpreted as the extent to which the same or different genes affect the two traits. Likelihood ratio tests were used to assess the evidence for deviations of the genetic correlation from zero and 1 (i.e., whether genes similarly or differentially affect the two traits). Detailed explanation of the bivariate model is outside the scope of this article but is available elsewhere [Almasy et al., 1999]. All genetic analyses were performed using the Sequential Oligogenic Linkage Analysis Routines (SOLAR) software [Almasy and Blangero, 1998]. General statistics, data manipulations, and plotting of figures were performed in the R statistical suite (R V2.8.1, R Foundation for Statistical Computing, Vienna, Austria).

The variance components methods described above assume normally distributed traits; however, caries scores exhibit skewed distributions. In general, variance components methods are known to be susceptible to large violations of normality [Allison et al., 1999]. Therefore, we addressed the possibility that model performance was adversely affected by deviations from normality by applying a severe normalization procedure to the data. First, caries scores were regressed on age and sex. Second, residuals were ranked and scaled (by dividing by the sample size + 1) to range between zero and one. Finally, we applied the probit function (the quantile function or inverse cumulative distribution function for the standard normal distribution) to scaled rankings in order to generate exactly normally distributed traits. Analysis of transformed traits yielded strikingly similar results (not shown), suggesting that the skewed distributions of untransformed traits were not noticeably reducing model performance.

Overall, power to detect heritability of 25% or greater was high. Power was estimated empirically in SOLAR via 1,000 simulations of a normally distributed, heritable trait. Power to detect heritability of $h^2 = 0.25$ was $>99\%$ for permanent caries scores, $81\%$ for primary caries scores, and $>99\%$ for total dentition caries scores. Power to detect heritability at $h^2 = 0.30$ was near 100% for permanent, primary, and total dentition caries scores.
Results

The demographic characteristics of the sample are shown in table 1. Prevalence of dental caries (i.e. the percentage of sample with one or more carious lesions) by age cohort is shown in table 2. As expected, prevalence of dental caries was greater in the primary dentition than permanent dentition for children (age groups <6 years and 6–14 years). For all age groups, and for both primary and permanent dentitions, the prevalence of caries was greater (or equal) for PFS compared to SMS.

Furthermore, mean proportion caries scores were greater for PFS than SMS (data not shown). For example, the mean proportions of carious PFS and SMS were 0.32 and 0.09, respectively, for permanent dentition, and 0.16 and 0.06, respectively, for primary dentition. Note that proportion caries scores (i.e. the proportion of carious surfaces out of the total at-risk surfaces) takes into account the fact that the full dentition includes many more SMS than PFS.

Distributions of PFS and SMS caries scores (i.e. dfs, DMFS, and proportions thereof) are shown in figure 1. These distributions further exemplified the differences in caries experience in PFS compared to SMS. While fewer individuals experienced SMS caries, some of those who did exhibited many affected SMS. Due to the fact that many more SMS than PFS are at-risk for caries, the dfs and DMFS scores for SMS show a greater range of values (and greater variance), and contributed substantially to the total disease burden in the sample. This observation was especially evident for caries occurrence in the permanent dentition.

Table 3 shows the heritability for PFS and SMS caries, as well as the genetic correlations between them, for primary, permanent, and total (i.e. all teeth present regardless of dentition type) dentitions. In general, caries scores were moderately to highly heritable (h$^2$ estimates ranged from 17 to 53%; p < 0.001 for all phenotypes), indicating that the cumulative effect of genes accounted for a substantial portion of disease risk. Moreover the proportional caries scores exhibited greater heritability than the corresponding dfs and DMFS scores, suggesting that proportional caries scores, which take into account both the number of affected surfaces and number of at-risk surfaces, may be improved metrics for studying the genetics of dental caries.

Heritability of caries scores was generally similar between PFS and SMS (i.e. estimates were within 1–2 standard errors of each other). These results indicate that similar proportions of phenotype variation are explained by genes for PFS and SMS caries.

<table>
<thead>
<tr>
<th>Demographic characteristic</th>
<th>Sample size</th>
<th>Number of kinships</th>
<th>Mean size of kinships (range)</th>
<th>Number of relative pairs</th>
<th>Parent-offspring</th>
<th>Siblings</th>
<th>Grandparent-grandchild</th>
<th>Half-siblings</th>
<th>Avuncular (i.e. uncle-nephew)</th>
<th>First cousins</th>
<th>Other relatives</th>
<th>Total related pairs</th>
<th>Within kinship unrelated pairs (i.e. spouses, etc.)</th>
<th>Total pairs</th>
<th>Self-reported whites, %</th>
<th>Females, %</th>
<th>Mean age, years (range)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sample size</td>
<td>2,663</td>
<td>740</td>
<td>4.72 (1–20)</td>
<td></td>
<td>1,736</td>
<td>676</td>
<td>60</td>
<td>322</td>
<td>124</td>
<td>98</td>
<td>35</td>
<td>3,051</td>
<td>739</td>
<td>3,790</td>
<td>89.64</td>
<td>55.61</td>
<td>19.83 (1–93)</td>
</tr>
</tbody>
</table>
statistically different from 1.0 for DMFS, dfs + DMFS, proportion dfs, proportion DMFS, and proportion dfs + DMFS. In contrast, the genetic correlation between PFS and SMS for the dfs caries score was 0.76, indicating that approximately 58% (i.e. $\rho G^2 = 0.76^2$) of the heritability of these two traits was explained by common genetic effects and approximately 42% by trait-specific genetic effects ($p < 0.001$). Trait-specific genetic effects may include genes that differentially impact dfs for PFS and SMS in magnitude and/or direction. We caution that while genetic correlation was not statistically different from 1.0 for all other caries scores, our results indicated a lack of evidence for surface type-specific genetic effects, rather than evidence for a lack of surface type-specific genetic effects. In other words, true surface type-specific genetic effects may have gone undetected, but if so, they were likely small in comparison to common genetic effects.
Overall, the results of this report were consistent with a previous study in this sample on tooth-level caries scores (i.e. d1ft, D1MFT, and d1ft + D1MFT in primary, permanent, and total dentitions, respectively) [Wang et al., 2010]. However, for all dentition types (i.e. primary, permanent, and total), heritability of both PFS and SMS were reduced compared to tooth-level caries scores. One explanation for the reduced heritability observed for PFS and SMS caries scores is that genetic factors may similarly affect caries risk for both surface types; therefore limiting analysis to one surface type may be less informative for the genetics of caries experience than simultaneous analysis of caries risk across both surface types. This notion is consistent with the high genetic correlation between PFS and SMS caries scores reported in this study.

While the current study did not seek to identify any specific genes affecting caries risk, it did show that PFS and SMS caries scores are heritable, and therefore may be useful phenotypes for future genetic association studies. The genetics of dental caries is thought to be very complex, with many biological mechanisms affecting cariogenesis and operating over long periods of time. Such mechanisms may include genes related to dietary choices (such as taste [Wendell et al., 2010] and olfactory receptors), immune response to pathogens [Bergandi et al., 2007], tooth enamel composition [Slayton et al., 2005], tooth morphology, saliva composition and flow rate, oral health behaviors, transmission of cariogenic bacteria among hosts [Law et al., 2007], and others. Genes acting through these and other avenues, some likely interacting with environmental factors, may each contribute only slightly to overall caries risk, and therefore teasing out the complex interplay of genetic and environmental risk factors represents a necessary challenge for better understanding of the multifactorial nature of dental caries.

Several limitations of the study warrant discussion. In particular, DMFS and dfs + DMFS caries scores for PFS and SMS were correlated, in part, due to the M surface classification. By convention, teeth missing due to decay contributed all surfaces to DMFS scores even though the actual carious lesion may have been limited to one surface type. Bias due to the M classification may have led to inflated estimation of the genetic correlation between PFS and SMS. In fact, the phenotypic correlation between PFS and SMS for the M component of DMFS was far greater than for D, F, or sound teeth (results not shown). However censoring the DMFS caries scores by removing the contribution of the M classification is also not ideal. Instead we suggest cautious interpretation of our genetic correlations given the known bias due to phenotype definitions. Likewise, restoration of approximal lesions is often performed via a two-surface filling, which results in the adjacent occlusal surface being counted toward PFS caries scores despite not necessarily having been decayed. This may have caused inflated PFS caries scores. Again, such bias may lead to inflated estimation of genetic correlation.

### Table 3. Heritability and genetic correlation of PFS and SMS caries scores

<table>
<thead>
<tr>
<th>Phenotype</th>
<th>n</th>
<th>PFS h²</th>
<th>SE</th>
<th>p²</th>
<th>R²</th>
<th>SMS h²</th>
<th>SE</th>
<th>p²</th>
<th>R²</th>
<th>Shared genetics</th>
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<td></td>
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<td></td>
<td></td>
<td></td>
<td>rhoG SE p rhoG=0 SE p rhoG=1 SE p</td>
</tr>
<tr>
<td>Primary dentition</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>dfs</td>
<td>953</td>
<td>0.34</td>
<td>0.10</td>
<td>3 × 10⁻⁴</td>
<td>0.01</td>
<td>0.42</td>
<td>0.10</td>
<td>2 × 10⁻⁵</td>
<td>0.00</td>
<td>0.76</td>
</tr>
<tr>
<td>Proportion dfs</td>
<td>953</td>
<td>0.53</td>
<td>0.11</td>
<td>2 × 10⁻⁶</td>
<td>0.04</td>
<td>0.42</td>
<td>0.10</td>
<td>2 × 10⁻⁵</td>
<td>0.00</td>
<td>0.93</td>
</tr>
<tr>
<td>Permanent dentition</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
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<td></td>
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</tr>
<tr>
<td>DMFS</td>
<td>1,939</td>
<td>0.27</td>
<td>0.06</td>
<td>7 × 10⁻⁶</td>
<td>0.35</td>
<td>0.27</td>
<td>0.07</td>
<td>2 × 10⁻⁴</td>
<td>0.22</td>
<td>1.00</td>
</tr>
<tr>
<td>Proportion DMFS</td>
<td>1,939</td>
<td>0.37</td>
<td>0.06</td>
<td>2 × 10⁻¹¹</td>
<td>0.33</td>
<td>0.21</td>
<td>0.06</td>
<td>2 × 10⁻⁴</td>
<td>0.21</td>
<td>0.98</td>
</tr>
<tr>
<td>Total dentition</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>dfs + DMFS</td>
<td>2,378</td>
<td>0.19</td>
<td>0.04</td>
<td>2 × 10⁻⁶</td>
<td>0.28</td>
<td>0.17</td>
<td>0.04</td>
<td>1 × 10⁻⁵</td>
<td>0.17</td>
<td>1.00</td>
</tr>
<tr>
<td>Proportion dfs + DMFS</td>
<td>2,378</td>
<td>0.31</td>
<td>0.04</td>
<td>2 × 10⁻¹³</td>
<td>0.35</td>
<td>0.23</td>
<td>0.04</td>
<td>2 × 10⁻⁸</td>
<td>0.19</td>
<td>0.99</td>
</tr>
</tbody>
</table>

h² = Residual heritability after adjusting for sex and age; SE = standard error, not applicable for models that converged at rhoG = 1.00; R² = proportion of trait variance attributable to sex and age; rhoG = genetic correlation, i.e. a measure of the extent to which traits covary due to genes.

a p value that PFS caries is heritable. b p value that SMS caries is heritable. c p value that shared genetic risk factors affect both PFS and SMS. d p value that unshared genetic risk factors differentially affect PFS and SMS.
Another limitation of this study was that all tooth surfaces were dichotomized as either PFS or SMS, which may not fully reflect the complex hierarchy of surface-specific cariologic risk factors. The epidemiology of surface-specific caries incidence suggests differences in caries susceptibility (and therefore suggests differences in caries risk factors) among PFS and among SMS [Batchelor and Sheiham, 2004; Psoter et al., 2009]. Future development of improved systems for categorizing and analyzing tooth surfaces may assist in discovering the surface-level factors leading to disease. For example, further subdivisions of surface types such as approximal versus buccal/lingual SMS, mandibular vs. maxillary, etc., may also be important for assessing the differential effects of environmental and genetic risk factors. Additionally, risk modification by age and, in particular, the duration of time that surfaces were present (i.e. at risk) is difficult to model for primary, permanent, and combined dentitions. Moreover, differences in environmental factors, such as fluoride varnish and sealants, were not modeled in these analyses, but may have differentially impacted caries scores among study participants leading to noise in the phenotype assessments. Indeed, other studies have shown important differential effects of fluoride exposure, dietary habits, socioeconomic status, and age on PFS and SMS caries [Maupome et al., 2001; Jiang et al., 2005; Warren et al., 2006].

Lastly, our method of data collection – visual inspection with dental explorer – is itself limited in that it is not the gold standard method for clinical caries assessment and may lead to deflated caries scores. However, assessment of data quality in our sample [Polk et al., 2008; Wendell et al., 2010] has shown that our method produces reliable and reproducible data of sufficient quality for pursuing the goals of this study. Moreover, our methods of data analysis are very robust. In general, measurement error or bias, model misspecification, and/or any unknown factors introducing noise into the phenotype would not lead to false positive findings (because measurement issues would not affect data proportional to biological relatedness among participants), but would instead bias these analyses toward the null hypothesis. Therefore, despite these limitations, the analyses are likely conservative and the general conclusions from this study are robust.

This study benefits from many strengths, including a large sample from an understudied, high-risk population. Our household-based recruitment yielded a rich sample comprised of a variety of relative types, which greatly improved our estimation of heritability and genetic correlation compared to studies limited to one or few relative types (i.e. studies of siblings, parent-offspring trios, etc.). This is because extended relatives (especially relatives residing in separate households such as half-siblings, cousins, etc.) are less likely to share familial nongenetic factors that could lead to overestimation. Our study was adequately powered, and all significant p values far exceeded the significance thresholds for Bonferroni adjustment for multiple comparisons. Additionally, we assessed the effect of nonnormality of our trait distributions by repeating analysis using exactly normal transformed phenotypes. Heritability and genetic correlation estimates were nearly identical, suggesting that deviations of caries scores from normality did not adversely affect our modeling framework.

The major conclusion from this study was that caries scores were heritable and that the majority of genes affecting caries risk were common to both PFS and SMS. While some surface type-specific genetic effects may exist, especially for dfs scores in primary dentition, the high genetic correlations suggest that combining PFS and SMS in future efforts to identify genes involved in caries risk may be beneficial complement to studying PFS and SMS separately. Additionally, this study highlighted the need for better caries phenotypes and/or modeling approaches that more accurately capture the complexity of the distribution of caries risk across the dentition. Additional work is currently needed to identify the environmental and genetic factors that differentially affect caries risk across surface types, and to devise better categorization methods that sensibly group tooth surfaces based on common risk profiles. This study is one of few attempts at defining new traits and genetic models which may assist in finding the specific genes implicated in caries etiology, and lead to improved understanding, and prevention, of the factors leading to disease.

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Community Hospital, Richwood, W.Va.; Summersville Memorial Hospital, Summersville, W.Va.; Webster County Memorial Hospital, Webster Springs, W.Va. Additional thanks is given to the GORGE Connection Rural Health Education Partnership Board, the Webster-Nicholas Rural Health Education Consortium Board, the West Virginia Rural Health Education Partnerships program, the Nicholas and Webster Boards of Education, the UPMC Braddock Community Advisory Board, and the individuals and social service agencies that helped develop the Center for Oral Health Research in Appalachia. Lastly, we would like to acknowledge three anonymous reviewers for their thoughtful feedback and commentary on this work.

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Co-author contributions are as follows: conceived and designed the COHRA initiative: R.J.W., R.C., D.W.N., M.L.M.; conceived and designed this study: J.R.S., X.W., R.J.W., K.T.C., M.L.M.; analyzed the data: J.R.S., X.W.; interpreted the results: J.R.S., X.W., R.S.D., S.W., R.J.W., K.T.C., R.C., D.W.M., M.L.M.; wrote the paper: J.R.S., X.W., R.J.W., K.T.C., M.L.M.

Disclosure Statement

All co-authors have no conflicts of interest.

References


