Effectiveness of the Air Purification Strategies for the Treatment of Allergic Asthma: A Meta-Analysis

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Keywords

Air purification · Allergy · Asthma · Meta-analysis

Abstract

We updated the meta-analysis published by McDonald et al. [Chest 2002;122:1535–1542] by reviewing the effectiveness of air purification for the treatment of home-related allergic asthma (dust mite, dog, cat, and cockroach). We analysed the trials included by McDonald et al. as well as studies published since 2000. Data on asthma symptoms scores (ASS), medication use, forced expiratory volume in 1 s as a percentage of the predicted value (FEV₁ %pred), histamine provocative concentration causing a 20% reduction in FEV₁ (PC₂₀), Asthma Quality of Life Questionnaire (AQLQ) scores, and fractional exhaled nitric oxide (FeNO) levels were extracted. The effectiveness was examined using metafor (registered in Prospero CRD42019127227). Ten trials including a total of 482 patients (baseline characteristics: mean FEV₁ %pred 83.2%, $I^2 = 96.7\%$; mean PC₂₀ 4.93 mg/mL, $I^2 = 44.0\%$; mean AQLQ 4.67 [max. 7], $I^2 = 93.7\%$; mean FeNO 36.5 ppb, $I^2 = 0\%$) were included. We assessed the mean differences in the AQLQ scores as +0.36 (95% CI 0.10 to 0.62, $p = 0.01$, $n = 302$,

$I^2 = 0\%$) and the FeNO levels as –6.67 ppb (95% CI –10.56 to –2.77, $p = 0.0008$, $n = 304$, $I^2 = 0\%$). The standardised mean differences in all other health outcomes were not significant (ASS –0.68, $p = 0.20$; medication use: –0.01, $p = 0.94$; FEV₁ %pred –0.11, $p = 0.34$; PC₂₀ +0.24, $p = 0.53$). We found statistically significant mean differences in the AQLQ scores and FeNO levels in patients with predominantly mild to moderate asthma at baseline. A large trial reported great improvement in the subgroup of patients receiving Global Initiative for Asthma step 4 therapy. We recommend that future studies on air purification focus on patients with severe and poorly controlled allergic asthma.

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Introduction

Respiratory allergy is a public health problem that affects approximately 400 million people [1]. The most common home-related respiratory allergies result from

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house dust mite, dog, cat, and cockroach allergen (Global Initiative for Asthma, GINA, 2018). Therapies such as pharmacological treatment, immunotherapy, and avoidance of indoor allergen exposure have been developed for the treatment of allergic asthma [2]. Evidence of clinical benefits of textile-based avoidance strategies has not been demonstrated in rigorous systematic reviews [3–5]. In a scoping review, Boven et al. [6] observed potential success with the strategy of air purification for the treatment of house dust mite allergy-related asthma. Previously, McDonald et al. [7] reported improvements in asthma symptom scores (ASS) associated with air purification in a small patient subgroup ($n = 88$).

Whether the purification of indoor air is of clinical importance in patients with asthma remains an unanswered question. An allergic reaction is provoked in the upper airways after the deposition of aerosol particles in the epithelium. The faecal pellets of house dust mites are very small in size, at 10–40 μm (mean 22 μm), and decrease when they are partially degraded over time (diameter >0.5 μm) [8, 9]. A large proportion of cat and dog allergens are smaller than 2 μm in diameter and coagulate in the air to other aerosol dust [10]. The particle size of cockroach allergens is mainly >10 μm [11]. Industrial branches have developed specific filters (high-efficiency particulate air, HEPA, filters) that capture very small airborne particles with high efficiency (at least 85–99.999995% of particles with a diameter of 0.3 μm) [12]. These HEPA filters are applied in residential products such as housing ventilation units, mobile air cleaners, nocturnal temperature-regulated laminar airflow units, and vacuum cleaners. The strategy of air purification has a potential advantage over a textile-based control strategy because the former strategy traps airborne allergens emitted from clothes as well as emissions from indoor textiles. This advantage may explain the clinical potential of the air purification strategy. As the current evidence on the clinical effectiveness of the air purification strategy is based on small sample sizes and was obtained many years ago, there is a need to update the evidence base, as new devices for purifying the nocturnal breathing zone have been introduced [13, 14].

This study updates the existing systematic review by McDonald et al. [7] entitled “Effect of Air Filtration Systems on Asthma” by reviewing the clinical effectiveness of the air purification strategy for the treatment of home-related allergic asthma (house dust mite allergy, dog allergy, cat allergy, and cockroach allergy).

Methods

Reference Search

The starting point of this study was the systematic review by McDonald et al. [7]. This meta-analysis included ten trials. An updated search of the literature published since January 2000 was performed in EMBASE, MEDLINE, and the Cochrane Central Register of Controlled Trials (CENTRAL). The trials were limited to peer-reviewed publications in the English language, and (Congress) abstracts were excluded from the analysis. The titles and/or abstracts of the studies retrieved during the search were screened (with Endnote) by the first author (F.E.v.B.) to identify randomised trials that met the inclusion criteria outlined below. The full texts of the potentially included trials were retrieved and assessed for inclusion by the first (F.E.v.B.) and second (N.W.d.J.) authors. Any ambiguities in the selections were resolved by discussion. The inclusion criteria were as follows:

- Type of study: randomised controlled trials with blinding.
- Intervention: housing or mobile ventilation systems, including HEPA filters but not vacuum cleaners.
- Participants: participants with physician-diagnosed bronchial allergic asthma. These participants had their sensitisation assessed by either skin testing or serum assays for specific IgE antibodies (house dust mite allergy, dog allergy, cat allergy, and cockroach allergy). The asthma assessment included a history of asthma symptoms and a pulmonary function test.
- Controls: participants who received a placebo or no treatment.

Data Extractions and Outcomes

The data were extracted by the first author (F.E.v.B.). The trials included in McDonald et al. [7] were re-extracted, as this review presented only the results but not the extracted data. The data extractions yielded the following: characteristics of the study population including the baseline data; type of intervention and the control; study methodology, and outcomes. Missing data were requested from the study authors. A second author (N.W.d.J.) verified the selections and the data extraction conducted by the first author. Any ambiguities in the selection and the extraction were resolved by discussion.

The main outcome(s) were: the asthma symptom score; the number of patients with improved outcomes; medication use; forced expiratory volume in 1 s as a percentage of the predicted value ($\text{FEV}_1\ \%\text{pred}$); provocative concentration that causes a 20% reduction in FEV_1 (PC_{20}); Asthma Quality of Life Questionnaire (AQLQ) score, and the fractional exhaled nitric oxide (FeNO) level. Additional outcomes included: the mite allergen load from the mattress ($\mu\text{g/g}$ dust); type of patient (child or adult), and the presence of primary and cosensitisation. These additional outcomes were all tested as possible explanatory variables in the presence of at least ten trials.

For the ASS, the PC_{20} , and the AQLQ scores, the final values were extracted (following Egbewale et al. [15]). The change scores were extracted for FEV_1 , medication use, and FeNO level. We defined the direction of changes as positive for an increasing FEV_1 and negative for a decreasing FeNO level and medication use.

Risk of Bias (Quality) Assessment

The risk of bias was assessed for the following domains: sequence generation, allocation concealment, blinding, incomplete outcome data, and selective outcome reporting. The assessment

Fig. 1. Flow chart of the reference search.

was performed by the first author (F.E.v.B.) with the Review Manager (RevMan) computer program version 5.3 (the Cochrane Collaboration, 2014; Nordic Cochrane Centre, Copenhagen, Denmark). A second author (N.W.d.J.) verified the assessments of the first author by considering a sample. Any ambiguities in the assessments were resolved by discussion.

Strategy for Data Synthesis

The effect size was set to the standardised mean difference, excluding the number of patients showing improvement (risk ratio). We chose the mean difference as the effect size in cases in which the outcomes were all measured in the same manner (AQLQ and FeNO). First, the overall effect of the health outcomes was estimated by a random-effects meta-analysis. Additionally, the I^2 was calculated for examining heterogeneity in the outcomes. In the absence of heterogeneity ($I^2 = 0$), a fixed-effects model was used. The explanatory variables of interest included the primary sensitisation (house dust mite allergy, dog allergy, cat allergy, or cockroach allergy), the mite allergen load from the mattress at baseline, possible confounding by the type of patient (child/adult), and the presence of cosensitisation. These outcomes were analysed for a preferred minimum of ten trials per variable [16]. All the calculations were performed with the metafor package in R [17, 18]. The level of significance was set to $\alpha = 0.05$.

Results

Selection of the References

We selected and included studies in two groups of publications. First, we screened the ten trials included in the meta-analysis by McDonald et al. [7]. Three trials

were excluded for a lack of or only partial reporting on the treatment of asthma [19, 20] or reporting incomplete data [21]. The remaining seven trials were included in the analysis [22–28].

The second group consisted of studies identified in our updated search (Fig. 1) [29]. We identified a total of 1,000 titles and abstracts. A total of 971 titles were excluded for lacking randomisation and/or blinding regarding the effectiveness of air purification. Twenty-nine potentially relevant titles were selected for inclusion. We excluded twenty-six full-text articles for not meeting our inclusion criteria (online suppl. Table; for all online suppl. material, see www.karger.com/doi/10.1159/000506284). Three full-text articles were included in the analysis [13, 30, 31]. In total, ten full-text articles were included in the meta-analysis.

Description of the Trials and the Baseline Characteristics

Ten trials published between 1973 and 2012 reported the treatment of asthma by air purification (Table 1). In four trials, the primary sensitisation was a pet allergy [13, 27, 28, 30]; five trials reported patients with house dust mite allergy [22–26], and one trial reported a mix of primary antigens [31]. None of the trials reported monosensitisation in the included patients. One trial [31] presented data on the specific IgE during the trial. Three trials reported the treatment of children with allergic asthma; the others reported the treatment of adults or both chil-

Table 1. Characteristics of the included studies

Trial	Use of a HEPA filter	Subjects	Primary allergy	Health outcomes extracted
Zwemer [22], 1973	Nocturnal laminar airflow	Child	House dust mite	ASS
Verrall [23], 1988	Nocturnal laminar airflow	Adult	House dust mite	Medication use
Antonicelli [23], 1991	Mobile device	Adult	House dust mite	ASS, medication use, FEV ₁ %pred, PC ₂₀
Warburton [24], 1994	Mobile device	Adult	House dust mite	FEV ₁ %pred
Van der Heide [26], 1997	Mobile device	Adult	House dust mite	PC ₂₀
Wood [28], 1998	Mobile device	Adult	Cat	ASS, medication use
Van der Heide [27], 1999	Mobile device	Child	Cat or dog	Medication use, PC ₂₀
Pedroletti [13], 2009	Nocturnal laminar airflow	Adult	Cat or dog	AQLQ score, FeNO level
Sulser [30], 2009	Mobile device	Adult	Cat or dog	PC ₂₀
Boyle [31], 2012	Nocturnal laminar airflow	Adult	House dust mite or cat	Medication use, FEV ₁ %pred, AQLQ score, FeNO level

	Random sequence generation (selection bias)	Allocation concealment (selection bias)	Blinding of participants and personnel (performance bias)	Incomplete outcome data (attrition bias)	Selective reporting (reporting bias)
Antonicelli, 1991	?	?	?	+	?
Boyle, 2012	+	+	+	+	+
Pedroletti, 2009	?	?	+	+	?
Sulser, 2009	?	?	+	+	?
Van der Heide, 1997	+	+	+	+	?
Van der Heide, 1999	+	+	+	+	?
Verrall, 1988	?	?	+	+	+
Warburton, 1994	?	?	+	+	?
Wood, 1998	?	?	+	+	?
Zwemer, 1973	?	?	+	+	?

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dren and adults. Four trials studied nocturnal laminar airflow in the breathing zone; the other six trials studied the use of a home ventilation or mobile device with a HEPA filter. Only one trial reported on the airborne allergen exposure [28], five other trials reported on dust exposure or allergen load at baseline [24–27, 30]. In the trial by Warburton et al. [25] only the data on FEV₁ %pred at baseline were available for analysis. In five trials, the mean FEV₁ %pred was 83.2% ($I^2 = 96.7\%$, $n = 346$). The mean PC₂₀ was 4.93 mg/mL ($I^2 = 44.0\%$, 2 trials, $n = 29$), the mean AQLQ score was 4.67 (max. 7; $I^2 = 93.7\%$, 2 trials, $n = 304$), and the mean FeNO level was 36.5 ppb ($I^2 = 0\%$, 2 trials, $n = 304$). For the ASS and medication use, we had no (quantitative) data available at baseline. Ten trials reported on the use of medication at baseline. In four trials, the change in the use of medication was a primary outcome for measuring effectiveness [22, 25, 26, 28]. Two investigations instructed their patients not to change their medication [23, 27]. In two trials [13, 31], patients were allowed to use more medication. The risk of bias was judged as predominantly unclear with a low risk of bias in blinding (Fig. 2).

Synthesis of the Efficacy Results

Four trials reported ASS as outcomes. We assessed the standardised mean difference in the ASS as -0.68 (95% CI -2.21 to 0.85 ; $p = 0.20$; $n = 77$; $I^2 = 51.0\%$; Fig. 3). The standardised mean difference in medication use was -0.01 (95% CI -0.22 to 0.21 ; $p = 0.94$; $n = 401$; $I^2 = 0\%$, 4 trials; Fig. 4). In three trials, the standardised mean difference in FEV₁ %pred was -0.11 (95% CI -0.34 to 0.12 ; $p = 0.34$; $n = 324$; $I^2 = 0\%$; Fig. 5). Four trials reported on the PC₂₀, with a standardised mean difference of $+0.24$ (95% CI -0.85 to 1.33 ; $p = 0.53$; $n = 98$; $I^2 =$

Fig. 2. Summary of the judgements on the risk of bias in the trials.

Fig. 3. Forest plot of the standardised mean differences in the ASS.

Fig. 4. Forest plot of the standardised mean differences in medication use.

Fig. 5. Forest plot of the standardised mean differences in the FEV₁ %pred.

60.0%; Fig. 6). The AQLQ scores were reported in two trials. We assessed the mean difference in the AQLQ scores as +0.36 (95% CI 0.10 to 0.62, $p = 0.01$, $n = 302$, $I^2 = 0\%$; Fig. 7). This positive increase was strongly influenced by the large trial by Boyle et al. [31] (weight 77%). The mean difference in the FeNO level was -6.67 ppb (95% CI -10.56 to -2.77, $p = 0.008$, $n = 304$, $I^2 = 0\%$; Fig. 8). None of the included trials reported on whether the physician-diagnosed numbers improved. Overall, the number of trials available was too small to allow any subgroup analysis.

Discussion

We reviewed the clinical effectiveness of the air purification strategy for the treatment of home-related allergic asthma in ten trials. The mean differences in the AQLQ score (MD = +0.36; $p = 0.01$) and the FeNO level (MD = -6.67; $p = 0.008$) were statistically significant, suggesting that asthma patients may benefit from air purification. These results were obtained in patients with predominantly mild to moderate asthma outcomes at baseline (the FEV₁ %pred, the AQLQ score, and the FeNO level).

Fig. 6. Forest plot of the standardised mean differences in the PC₂₀.

Fig. 7. Forest plot of the mean differences in the AQLQ scores.

Fig. 8. Forest plot of the mean differences in the FeNO levels.

The overall airway hyperresponsiveness was mild at baseline, according to the classification by Cockcroft et al. [32]. The risk of bias in the trials was predominantly judged unclear; however, blinding has a low risk of bias.

The strength of this meta-analysis was the rigorous selection of trials and extraction of data. We decided a priori whether to extract change or final values considering the statistical notes by Egbevale et al. [15]. In our study, we excluded some trials that were included by McDonald et al. [7] due to a critical process in extracting the data. For instance, they included the ASS by Reisman et al. [20]. After a critical review of this paper, we decided not to extract these data as only 11 of 32 patients were diagnosed

with asthma; thus, we excluded this trial from the analysis. We noticed that this trial was also excluded for the same reason in the meta-analysis by Gøtzsche and Johansen [3]. While the previously analysed trials were quite old, the recent trials included the use of validated outcomes such as the AQLQ score [33]. In patients with mild to moderate disease, we observed small (not reaching the minimum clinically important difference) but significant improvements in the AQLQ scores and FeNO levels. This effect could possibly be stronger in patients with severe asthma than in those with mild to moderate asthma. This possible tendency is well presented in the large trial by Boyle et al. [31]. They studied the effectiveness of the Pro-

texo system (a nocturnal temperature-controlled laminar airflow) and reported the outcomes of the use of medication, FEV₁ %pred, AQLQ scores, and FeNO levels. They differentiated the AQLQ score, their primary outcome, and the asthma status defined by the treatment intensity of GINA and the asthma control test (ACT). After a 1-year treatment period, Boyle et al. [31] reported an AQLQ score difference of +0.31 ($p = 0.04$) in all the studied patients ($n = 282$). When limited to the patients classified as requiring GINA step 4 therapy (GINA 4) at baseline, the difference became +0.47 ($p = 0.04$, $n = 129$). In the patients receiving GINA 4 with poor control (ACT <18), the difference in the AQLQ score was +0.70 ($p = 0.02$, $n = 87$). Additionally, in the patients with a high FeNO level at baseline, the same tendency was reported by Boyle et al. [31] (mean difference in FeNO -29.7 ppb, $p = 0.001$).

The limitation of this meta-analysis was the relatively small number of trials included in the analysis. Our update did not result in many new included trials. In total, we included the same number of trials ($n = 10$) as McDonald et al. [7] included in their earlier meta-analysis. We had to exclude three trials that were included by McDonald et al. [7] because of a lack of reporting on the treatment of asthma or incompleteness of the data. McDonald et al. [7] previously reported “a small but statistically significant difference in total symptoms associated with use of domestic air filters.” They did not find benefits associated with medication use or morning peak flow values. In our update, we did not find a significant difference for the ASS outcome. The significance reported by McDonald et al. [7] was based on an analysis by the fixed-effects model. As the ASS showed moderate heterogeneity ($I^2 = 51\%$), we introduced the random-effects model and the significance was lost. The use of domestic HEPA filters will also be of relevance in the treatment of non-allergic asthma, for instance by filtering indoor air pollution. As we included only trials on the treatment of allergic asthma, this possible issue did not bias our results. The description of the allergen exposure differed in the trials and was sometimes poorly presented. Therefore, we could not analyse the degree of the exposure, and also cannot exclude the possibility that a variation of allergens from other sources affected the results.

The significant differences we found were both a result of trials sponsored by Airsonett AB (Ängelholm, Sweden). One of these trials [31] was predominantly responsible for the positive AQLQ score analysis and was judged as having a risk of bias in randomisation. Their treatment group was twice the size of the control group. In principle, this creates a risk of selection bias as recruiters could “guess with greater than a 50% probability what the next

treatment allocation will be” [34]. In their report, we did not find indications for baseline imbalances biasing the estimates. Another issue of relevance in both trials on the Protexo system is the possibility of changes in medication use. Pedroletti et al. [13] reported that “inhaled, short-acting beta-2 agonists were allowed as rescue treatment.” Boyle et al. [31] instructed that the patients “asthma medication were kept unchanged for the first 3 months, and thereafter adjusted to optimise asthma control.” We cannot exclude the possibility that these instructions confounded the significant results we found. Overall, the results require independent repeating, with careful monitoring of allergen exposure.

Other studies on the Protexo system resulted in (some) clinical benefits. Schauer et al. [35] observed reduced asthma exacerbation and hospitalisations in an observational study in patients with predominantly difficult-to-control asthma. In a recent pilot study, Gore et al. [36] reported the potential for the use of the Protexo system as an add-on to standard pharmacological treatment in children with difficult-to-control atopic dermatitis. These results also reflect the need to study patients with severe and uncontrolled conditions.

In brief, we reviewed the clinical effectiveness of the air purification strategy for the treatment of home-related allergic asthma (house dust mite allergy, dog allergy, cat allergy, and cockroach allergy). We found statistically significant mean differences in the AQLQ scores and FeNO levels in patients with predominantly mild to moderate asthma at baseline. A large underlying trial [31] showed potentially great improvement in the AQLQ scores in the subgroup of patients receiving GINA 4 therapy with poor control. Future studies on air purification strategies with rigorous trial designs that focus on patients with severe and poorly controlled allergic asthma are warranted.

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Statement of Ethics

Ethical approval for this meta-analysis was not required.

Disclosure Statement

The authors have no conflicts of interest to declare.

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Author Contributions

All authors contributed to the design of the work. F.E.v.B. and N.W.d.J. selected the references and extracted the data. F.E.v.B. and L.R.A. analysed the data. G.J.B. and R.G.v.W. contributed to the interpretation of the data. All authors contributed to the draft of the work, and read and approved the final manuscript.

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